

LLDA Variational Approximation Derivation

Model Parameters

$$\begin{aligned}
 \beta &= P(S|Z) \\
 \pi &= P(L|Z) \\
 \alpha &= P(\theta) \\
 \theta &= P(Z|\theta)
 \end{aligned} \tag{1}$$

Variational Parameters

Where D is the experiment.

$$\begin{aligned}
 \gamma &= P(\theta|D) \\
 \phi &= P(Z|S, D) \\
 \psi &= P(L|S, D)
 \end{aligned} \tag{2}$$

Node Distributions

$$p(\theta; \alpha) = \frac{\Gamma(\sum_{i=1}^K \alpha_i)}{\prod_{i=1}^K \Gamma(\alpha_i)} \prod_{i=1}^K \theta_i^{\alpha_i - 1} \tag{3}$$

$$p(Z = z|\theta) = \theta_z \tag{4}$$

$$p(S = s|Z = z; \beta) = \beta_{sz} \tag{5}$$

$$p(L = l|Z = z; \pi) = \pi_{lz} \tag{6}$$

Likelihood

$$p(\theta, Z, S, L|\alpha, \beta, \pi) = p(\theta; \alpha) \prod_{n=1}^N p(Z_n|\theta)p(S_n|Z_n; \beta)p(L_n|Z_n; \pi) \tag{7}$$

$$p(S, L|\alpha, \beta, \pi) = \int_{\theta} p(\theta; \alpha) \left(\prod_{n=1}^N \sum_{Z_n=1}^K p(Z_n|\theta)p(S_n|Z_n; \beta)p(L_n|Z_n; \pi) \right) d\theta \tag{8}$$

We will use the notation \mathcal{D} to indicate the corpus of experiments. \mathcal{D} is a $V \times M$ matrix where there are V genes and M experiments in the corpus.

The log-likelihood function becomes

$$ll(\alpha, \beta, \pi) = \sum_{d=1}^M \log p(s_d, l_d|\alpha, \beta, \pi). \tag{9}$$

Multiplying and dividing by a variational distribution q gives

$$ll(\alpha, \beta, \pi) = \log \int_{\theta} \sum_{\underline{z}} \sum_{\underline{l}} p(\theta, Z, S, L|\alpha, \beta, \pi) \frac{q(\theta, Z, L)}{q(\theta, Z, L)} d\theta. \tag{10}$$

We have chosen to marginalize the label in this equation because it is only a partially observed random variable. In the cases the label is observed for a gene, the average becomes degenerate.

Applying Jensen's Inequality gives

$$ll(\alpha, \beta, \pi) \geq E_q[\log p(\theta, Z, L, S|\alpha, \beta, \pi)] - E_q[\log q(\theta, Z, L)]. \tag{11}$$

Choosing the fully factorized distribution for q as a variational distribution gives

$$q(\theta, Z, L|\gamma, \phi, \psi) = q(\theta; \gamma)q(Z; \phi)q(L; \psi). \tag{12}$$

Then the bound to the log-likelihood in 11 can be factorized to give

$$\begin{aligned} \mathcal{L}(\gamma, \phi, \psi; \alpha, \beta, \pi) = & \\ & E_q[p(\theta; \alpha)] + E_q[\log p(Z|\theta)] + E_q[p(S|Z; \beta)] + E_q[p(L|Z; \pi)] \\ & - E_q[\log q(\theta; \gamma)] - E_q[\log q(Z; \phi)] - E_q[\log q(L; \psi)] \end{aligned} \quad (13)$$

Variational Bound Expectations

We now compute each of expectations in equation 13.

$$E_q[p(\theta|\alpha)] = \log \Gamma \left(\sum_i \alpha_i \right) - \sum_i \log \Gamma(\alpha_i) + \sum_{i=1}^K (\alpha_i - 1) \left[\Psi(\gamma_i) - \Psi \left(\sum_j \gamma_j \right) \right] \quad (14)$$

$$E_q[p(Z = i|\theta)] = \sum_{n=1}^N \sum_{i=1}^K \phi_{ni} \left[\Psi(\gamma_i) - \Psi \left(\sum_j \gamma_j \right) \right] \quad (15)$$

$$E_q[p(S = v|Z = i; \beta)] = \sum_{n=1}^N \sum_{k=1}^V \phi_{ni} s_n^k \log \beta_{ki} \quad (16)$$

$$E_q[p(L = j|Z = i; \pi)] = \sum_{n=1}^N \sum_{i=1}^K \sum_{j=1}^L \phi_{ni} \psi_{nj} \log \pi_{ji} \quad (17)$$

$$E_q[\log q(\theta)] = \log \Gamma \left(\sum_i \gamma_i \right) - \sum_i \log \Gamma(\gamma_i) + \sum_{i=1}^K (\gamma_i - 1) \left[\Psi(\gamma_i) - \Psi \left(\sum_j \gamma_j \right) \right] \quad (18)$$

$$E_q[\log q(Z; \phi)] = \sum_{n=1}^N \sum_{i=1}^K \phi_{ni} \log \phi_{ni} \quad (19)$$

$$E_q[\log q(L; \psi)] = \sum_{n=1}^N \sum_{j=1}^L \psi_{nj} \log \psi_{nj} \quad (20)$$

Variational Parameter Estimation

Estimate of ϕ Isolating the terms in equation 13 that reference ϕ and including the constraint that $\sum_i \phi_{ni} = 1, \forall n$ gives

$$\mathcal{L}_{[\phi]} = \phi_{ni} \left[\Psi(\gamma_i) - \Psi \left(\sum_j \gamma_j \right) \right] + \sum_{k=1}^V \phi_{ni} s_n^k \log \beta_{ki} + \sum_{j=1}^L \phi_{ni} \psi_{nj} \log \pi_{ji} - \phi_{ni} \log \phi_{ni} + \lambda_n \left(\sum_{j=1}^K \phi_{nj} - 1 \right). \quad (21)$$

Differentiating with respect to the parameter and setting equal to 0 gives

$$\phi_{ni} \propto \exp \left(\Psi(\gamma_i) + \sum_{j=1}^L \psi_{nj} \log \pi_{ji} + \log \beta_{vi} \right) \quad (22)$$

Estimate of ψ Isolating the terms in equation 13 that reference ψ and including the constraint that $\sum_i \psi_{ni} = 1, \forall n$ gives

$$\mathcal{L}_{[\psi]} = \sum_{i=1}^K \phi_{ni} \psi_{nj} \log \pi_{ji} - \psi_{nj} \log \psi_{nj} + \lambda_n \left(\sum_{j=1}^L \psi_{nj} - 1 \right). \quad (23)$$

Differentiating with respect to the parameter and setting equal to 0 gives

$$\psi_{nj} \propto \exp \left(\sum_{i=1}^K \phi_{ni} \log \pi_{ji} \right) \quad (24)$$

Estimate of γ Isolating the terms in equation 13 that reference γ gives

$$\mathcal{L}_{[\gamma]} = (\alpha_i - 1) \left[\Psi(\gamma_i) - \Psi\left(\sum_j \gamma_j\right) \right] + \sum_{n=1}^N \phi_{ni} \left[\Psi(\gamma_i) - \Psi\left(\sum_j \gamma_j\right) \right]. \quad (25)$$

Differentiating with respect to the parameter and setting equal to 0 gives

$$\gamma_i = \alpha_i + \sum_{n=1}^N \phi_{ni} \quad (26)$$

Model Parameter Estimation

Estimate of β

$$\mathcal{L}_{[\beta]} = \sum_{d=1}^M \sum_{n=1}^{N_d} \phi_{ni}^d s_{dn}^v \log \beta_{vi} + \lambda_i \left(\sum_{k=1}^V \beta_{vi} - 1 \right) \quad (27)$$

The update equation becomes

$$\beta_{vi} \propto \sum_{d=1}^M \sum_{n=1}^{N_d} \phi_{ni}^d s_{dn}^v \quad (28)$$

Estimate of π

$$\mathcal{L}_{[\pi]} = \sum_{d=1}^M \sum_{n=1}^{N_d} \phi_{ni} \psi_{nj} \log \pi_{ji} + \lambda_i \left(\sum_{j=1}^L \pi_{ji} - 1 \right) \quad (29)$$

The update equation becomes

$$\pi_{ji} \propto \sum_{d=1}^M \sum_{n=1}^{N_d} \phi_{ni} \psi_{nj} \quad (30)$$

Estimate of α Since the parameters in the α vector are coupled a gradient descent procedure must be used. See [Blei et. al. 2002] for details.